

## PALM DISTRIBUTION AND LIMIT THEOREMS FOR RANDOM POINT PROCESSES IN $R^n$

BY

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**1. Introduction.** The paper deals with homogeneous random point processes in  $R^n$ . Our aim is to obtain sufficient conditions for asymptotic normality of the number of points of a point process  $P$  in a ball when the radius tends to infinity. These conditions are formulated in terms of Palm distributions of the point process  $P$ . The proofs of the theorems stated below are based on the relation between the distribution  $P$  of a point process in  $R^n$  and its Palm distribution due to Ambartzumian [1]. In the one-dimensional case these relations reduce to the so-called Palm-Khinchin formulae [2].

**2. Notation.** Let  $M$  be the class of all countable subsets of  $R^n$  such that any  $m \in M$  has no cluster point in a bounded subset of  $R^n$ .

Define  $N(B, m)$  to be the number of points in  $B \cap m$ , where  $B$  is a bounded Borel set in  $R^n$  and  $m \in M$ .

Denote by  $C$  the minimal  $\sigma$ -algebra of subsets of  $M$  containing all subsets of the form  $\{m: N(B, m) = k\}$ ,  $k = 0, 1, 2, \dots$ . Any probability measure  $P$  on  $C$  describes a random point process.

A random point process  $P$  is said to be *homogeneous* if, for any  $c \in C$ ,  $P(tc)$  does not depend on  $t \in T$ , where  $T$  denotes the group of all translations of  $R^n$  and

$$tc = \{m: t^{-1}m \in c\}.$$

Further, we assume that for every bounded Borel set  $B$  in  $R^n$

$$(1) \quad E_p(N(B)) = \lambda |B|, \quad \lambda < \infty,$$

where  $|B|$  is the volume of  $B$ . In other words, we consider the finite intensity case.

3. **Main results.** Let  $S(v)$  be the sphere of volume  $v$  centred at the origin in  $R^n$ . We consider the random number  $N(S(v))$  of points of the process in  $S(v)$  for large values of  $v$ .

$N(S(v))$  is called *asymptotically normal* if

$$(2) \quad \sum_{k: (k-\lambda v)/\sqrt{\lambda v} < \alpha} P(N(S(v)) = k) \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} \exp\left\{-\frac{s^2}{2}\right\} ds \quad \text{as } v \rightarrow \infty, \alpha \in R^1.$$

$N(S(v))$  is called *locally asymptotically normal in variation* if

$$(3) \quad \sum_{k=0}^{\infty} \left| P(N(S(v)) = k) - \frac{1}{\sqrt{2\pi\lambda v}} \exp\left\{-\frac{(k-\lambda v)^2}{2\lambda v}\right\} \right| \rightarrow 0 \quad \text{as } v \rightarrow \infty.$$

$N(S(v))$  is called *locally asymptotically normal* if

$$(4) \quad \sup_k \left| \sqrt{\lambda v} P(N(S(v)) = k) - \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(k-\lambda v)^2}{2\lambda v}\right\} \right| \rightarrow 0 \quad \text{as } v \rightarrow \infty.$$

Note that each of the relations (3) and (4) implies (2). It is known that under the assumption (1) the limit

$$(5) \quad \pi_k(v) = \lim_{h \rightarrow 0} \frac{P(\{N(S(v)) = k\} \cap \{N(S(v+h) - S(v)) = 1\})}{P(N(S(v+h) - S(v)) = 1)}$$

exists [1].

We call this limit the *spherical Palm distribution*. This limit can be interpreted as the conditional probability of  $\{N(S(v)) = k\}$  under the condition that a point of the random process  $P$  lies on the boundary of the sphere  $S(v)$ . By the *Palm distribution* we usually mean the conditional probability of  $\{N(S(v)) = k\}$  under the condition that there is a point of the process  $P$  at the origin. The spherical Palm distribution can be found by integration of the usual Palm distribution over the boundary of the sphere  $S(v)$  (see [1]).

The conditions for asymptotic normality will be given in terms of the variational distance

$$(6) \quad \varrho(v) = \sum_{k=0}^{\infty} |P_k(v) - \pi_k(v)|, \quad P_k(v) = P(N(S(v)) = k).$$

We show that the above-mentioned types of asymptotic normality of  $N(S(v))$  are implied by various assumptions concerning the rate of convergence of  $\varrho(v)$  to zero as  $v$  tends to infinity. In this sense the condition on the rate of convergence of  $\varrho(v)$  to zero can replace the usual mixing condition [3].

In the sequel we prove the following theorems:

THEOREM 1. If

$$\lim_{v \rightarrow \infty} \frac{1}{\sqrt{v}} \int_0^v \varrho(u) du = 0,$$

then  $N(S(v))$  is asymptotically normal.

THEOREM 2. If

$$\int_0^{\infty} \varrho(u) du < \infty,$$

then  $N(S(v))$  is locally asymptotically normal in variation.

THEOREM 3. If

$$\lim_{v \rightarrow \infty} \frac{1}{\sqrt{v}} \int_0^v u \varrho(u) du = 0,$$

then  $N(S(v))$  is locally asymptotically normal.

The proofs of these theorems are based on the following Ambartzumian relations (see [1]):

$$(7) \quad \begin{aligned} \frac{dP_0(v)}{dv} &= -\lambda \pi_0(v), \\ \frac{dP_k(v)}{dv} &= -\lambda (\pi_k(v) - \pi_{k-1}(v)), \quad k = 1, 2, \dots, \\ P_k(0) &= \begin{cases} 1 & \text{for } k = 0, \\ 0 & \text{for } k \neq 0. \end{cases} \end{aligned}$$

Further, without loss of generality we can assume that  $\lambda = 1$ .

In the one-dimensional case, Theorems 1 and 2 were announced in [4].

The role of the Palm distribution in the problems related to asymptotic normality of  $N(S(v))$  was first noticed by R. V. Ambartzumian to whom Theorem 1 should be attributed.

The author expresses his gratitude to Professor R. V. Ambartzumian for suggesting the present topic.

4. Proof of Theorem 1. Rewrite (7) in the form

$$(8) \quad \begin{aligned} \frac{dP_0(v)}{dv} &= -P_0(v) + (P_0(v) - \pi_0(v)), \\ \frac{dP_k(v)}{dv} &= -(P_k(v) - P_{k-1}(v)) + (P_k(v) - \pi_k(v)) + (\pi_{k-1}(v) - P_{k-1}(v)), \\ & \quad k = 1, 2, \dots \end{aligned}$$

We introduce the generating functions

$$\Pi_z(v) = \sum_{k=0}^{\infty} P_k(v) z^k, \quad A_z(v) = \sum_{k=0}^{\infty} (\pi_k(v) - P_k(v)) z^k$$

for which from (8) we derive

$$(9) \quad \frac{d\Pi_z(v)}{dv} = (z-1)\Pi_z(v) + (z-1)A_z(v), \quad \Pi_z(0) = 1.$$

Hence we get

$$(10) \quad \Pi_z(v) = e^{(z-1)v} + (z-1)e^{(z-1)v} \int_0^v e^{(1-z)u} A_z(u) du.$$

The characteristic function of the distribution  $\{P_k(v)\}$  can be obtained by substituting  $z$  by  $e^{it}$ . Hence it is enough to show that for fixed  $t$

$$\exp\{-it\sqrt{v}\} \Pi_{\exp(it/\sqrt{v})}(v) \rightarrow \exp\{-t^2/2\} \quad \text{as } v \rightarrow \infty.$$

Since the Poisson distribution is asymptotically normal, it remains to show that the contribution of the second summand in (10) vanishes as  $v$  tends to infinity.

We have

$$\begin{aligned} & \left| \exp\{-it\sqrt{v}\} \left( \exp\left\{\frac{it}{\sqrt{v}}\right\} - 1 \right) \times \right. \\ & \quad \times \int_0^v \exp\left\{ \left[ \exp\left(\frac{it}{\sqrt{v}}\right) - 1 \right] (v-u) \right\} A_{\exp(it/\sqrt{v})}(u) du \Big| \\ & \leq \left| \exp\left\{\frac{it}{\sqrt{v}}\right\} - 1 \right| \int_0^v \left| \exp\left\{ \left[ \exp\left(\frac{it}{\sqrt{v}}\right) - 1 \right] (v-u) \right\} \right| |A_{\exp(it/\sqrt{v})}(u)| du \\ & \leq \frac{t}{\sqrt{v}} \int_0^v \exp\left\{ \left( \cos\frac{t}{\sqrt{v}} - 1 \right) (v-u) \right\} |A_{\exp(it/\sqrt{v})}(u)| du. \end{aligned}$$

Since

$[\cos(t/\sqrt{v}) - 1](v-u) \leq 0$  for  $u \in (0, v)$  and  $|A_{\exp(it/\sqrt{v})}(u)| \leq \varrho(u)$ , the last expression does not exceed

$$\frac{t}{\sqrt{v}} \int_0^v \varrho(u) du.$$

Hence Theorem 1 holds.

**5. Proof of Theorem 2.** We first show that under the assumptions of Theorem 2 we have

$$\sum_{k=0}^{\infty} |P_k(v) - A_k(v)| \rightarrow 0 \quad \text{as } v \rightarrow \infty,$$

where

$$A_k(v) = e^{-v} \frac{v^k}{k!}, \quad k = 0, 1, 2, \dots$$

Since the probabilities  $A_k(v)$  satisfy the equations

$$(11) \quad \begin{aligned} \frac{dA_0(v)}{dv} &= -A_0(v), \\ \frac{dA_k(v)}{dv} &= -(A_k(v) - A_{k-1}(v)), \quad k = 1, 2, \dots, \\ A_k(0) &= \begin{cases} 1 & \text{for } k = 0, \\ 0 & \text{for } k \neq 0, \end{cases} \end{aligned}$$

using (7) we can write

$$(12) \quad \begin{aligned} \frac{d}{dv} (P_0(v) - A_0(v)) &= -(P_0(v) - A_0(v)) - (\pi_0(v) - P_0(v)), \\ \frac{d}{dv} (P_k(v) - A_k(v)) &= (P_{k-1}(v) - A_{k-1}(v)) - (P_k(v) - A_k(v)) + \\ &\quad + (\pi_{k-1}(v) - P_{k-1}(v)) - (\pi_k(v) - P_k(v)), \quad k = 1, 2, \dots \end{aligned}$$

Putting

$$\alpha_k(v) = P_k(v) - A_k(v), \quad \beta_k(v) = \pi_k(v) - P_k(v), \quad k = 0, 1, 2, \dots$$

we introduce the generating functions

$$A_z(v) = \sum_{k=0}^{\infty} \alpha_k(v) z^k, \quad B_z(v) = \sum_{k=0}^{\infty} \beta_k(v) z^k.$$

Using (12) we obtain the differential equation

$$(13) \quad \frac{dA_z(v)}{dv} = (z-1)A_z(v) + (z-1)B_z(v), \quad A_z(0) = 0.$$

Resolving this equation we have

$$(14) \quad A_z(v) = (z-1) e^{(z-1)v} \int_0^v e^{-(z-1)u} B_z(u) du.$$

Let  $D$  stand for the differentiation operation with respect to  $z$ . Since

$$\frac{1}{k!} D^{(k)} A_z(v)|_{z=0} = P_k(v) - A_k(v),$$

we obtain

$$\sum_{k=0}^{\infty} |P_k(v) - A_k(v)| = \sum_{k=0}^{\infty} \frac{1}{k!} |D^{(k)} A_z(v)|_{z=0}.$$

Further, we get

$$\begin{aligned} D^{(k)} A_z(v) &= D^{(k)} [(z-1)e^{(z-1)v} \int_0^v e^{(1-z)u} B_z(u) du] \\ &= e^{-v} \int_0^v e^u D^{(k)} [e^{z(v-u)} (z-1) B_z(u)] du. \end{aligned}$$

Using the formula

$$D^{(k)} e^{\lambda z} u(z) = e^{\lambda z} (D + \lambda)^{(k)} u(z),$$

we can write

$$\begin{aligned} D^{(k)} A_z(v) &= e^{-v} \int_0^v e^u e^{z(v-u)} (D + v - u)^{(k)} [(z-1) B_z(u)] du \\ &= e^{-v} \int_0^v e^u e^{z(v-u)} \sum_{j=0}^k \frac{k!}{j!(k-j)!} (v-u)^j D^{(k-j)} [(z-1) B_z(u)] du. \end{aligned}$$

Hence

$$(15) \quad \left. \frac{D^{(k)} A_z(v)}{k!} \right|_{z=0} = e^{-v} \int_0^v e^u \sum_{j=0}^k \frac{(v-u)^j}{j!} \left. \frac{D^{(k-j)} (z-1) B_z(u)}{(k-j)!} \right|_{z=0} du.$$

Further, we obtain

$$(16) \quad \begin{aligned} D^{(k-j)} (z-1) B_z(u)|_{z=0} &= D^{(k-j)} z B_z(u)|_{z=0} - D^{(k-j)} B_z(u)|_{z=0} \\ &= (k-j)! (\beta_{k-j-1}(u) - \beta_{k-j}(u)), \quad k = 0, 1, 2, \dots, j = 0, 1, \dots, k, \end{aligned}$$

where we put  $\beta_{-1}(u) \equiv 0$ . Substituting (16) in (15) we have

$$\left. \frac{D^{(k)} A_z(v)}{k!} \right|_{z=0} = e^{-v} \int_0^v e^u \sum_{j=0}^k \frac{(v-u)^j}{j!} (\beta_{k-j-1}(u) - \beta_{k-j}(u)) du.$$

Hence

$$\begin{aligned} \sum_{k=0}^{\infty} |P_k(v) - A_k(v)| &= \sum_{k=0}^{\infty} \left| \left. \frac{D^{(k)} A_z(v)}{k!} \right|_{z=0} \right| \\ &= e^{-v} \sum_{k=0}^{\infty} \left| \int_0^v e^u \sum_{j=0}^k \frac{(v-u)^j}{j!} (\beta_{k-j}(u) - \beta_{k-j-1}(u)) du \right|. \end{aligned}$$

We write

$$\sum_{j=0}^k \frac{(v-u)^j}{j!} (\beta_{k-j}(u) - \beta_{k-j-1}(u)) = \sum_{j=0}^k \left( \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right) \beta_{k-j}(u),$$

where, by definition,  $(v-u)^{-1}/(-1)! \equiv 0$ . Therefore, we get

$$\begin{aligned} \sum_{k=0}^{\infty} |P_k(v) - A_k(v)| &= e^{-v} \sum_{k=0}^{\infty} \left| \int_0^v e^u \sum_{j=0}^k \left( \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right) \beta_{k-j}(u) du \right| \\ &\leq e^{-v} \sum_{k=0}^{\infty} \int_0^v e^u \sum_{j=0}^k \left| \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right| |\beta_{k-j}(u)| du \\ &\leq e^{-v} \int_0^{v_0} e^u \sum_{j=0}^{\infty} \left| \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right| \varrho(u) du \\ &= e^{-v} \int_0^{v_0} e^u \sum_{j=0}^{\infty} \left| \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right| \varrho(u) du + \\ &+ e^{-v} \int_{v_0}^v e^u \sum_{j=0}^{\infty} \left| \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right| \varrho(u) du, \quad 0 < v_0 < v. \end{aligned}$$

Consequently, we obtain

$$e^{-v} \int_{v_0}^v e^u \sum_{j=0}^{\infty} \left| \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right| \varrho(u) du \leq 2 \int_{v_0}^{\infty} \varrho(u) du$$

and

$$\begin{aligned} &e^{-v} \int_0^{v_0} e^u \sum_{j=0}^{\infty} \left| \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right| \varrho(u) du \\ &= e^{-v} \int_0^{v_0} e^u \sum_{j=0}^{[v-u]} \left( \frac{(v-u)^j}{j!} - \frac{(v-u)^{j-1}}{(j-1)!} \right) \varrho(u) du + \\ &+ e^{-v} \int_0^{v_0} e^u \sum_{j=[v-u]+1}^{\infty} \left( \frac{(v-u)^{j-1}}{(j-1)!} - \frac{(v-u)^j}{j!} \right) \varrho(u) du \\ &\leq c \int_0^{v_0} e^{u-v} \frac{(v-u)^{[v-u]}}{[v-u]!} \varrho(u) du \leq \frac{\hat{c}}{\sqrt{[v-v_0]}} \int_0^{\infty} \varrho(u) du, \quad 0 < c, \hat{c} < \infty. \end{aligned}$$

In the last inequality we applied Stirling's formula. Finally, we get

$$\sum_{k=0}^{\infty} |P_k(v) - A_k(v)| \leq \int_{v_0}^{\infty} \varrho(u) du + \frac{\hat{c}}{\sqrt{[v-v_0]}} \int_0^{\infty} \varrho(u) du.$$

By choosing  $v_0$  and  $v$  sufficiently large the last expression can be made arbitrarily small. Since Theorem 2 is true for the Poisson distribution (see [5]), the proof is complete.

**6. Proof of Theorem 3.** It is sufficient to show that

$$(17) \quad \sup_k |\sqrt{v} P_k(v) - \sqrt{v} A_k(v)| \rightarrow 0 \quad \text{as } v \rightarrow \infty.$$

By (17) and the local limit theorem for the Poisson distribution, Theorem 3 holds.

By the converse formula for the Fourier transformation of the sequences  $\alpha_k(v)$  and (14) we have

$$\begin{aligned} P_k(v) - A_k(v) &= \frac{1}{2\pi} \int_{-\pi}^{\pi} A_{\exp(it)}(v) e^{-itk} dt \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} (e^{it} - 1) \exp\{v(e^{it} - 1)\} \left[ \int_0^v \exp\{(1 - e^{it})u\} B_{\exp(it)}(u) du \right] dt. \end{aligned}$$

Replacing  $t$  by  $t/\sqrt{v}$ , we get

$$\begin{aligned} \sqrt{v} P_k(v) - \sqrt{v} A_k(v) &= \frac{1}{2\pi} \int_{-\pi\sqrt{v}}^{\pi\sqrt{v}} \left( \exp\left\{ \frac{it}{\sqrt{v}} \right\} - 1 \right) \times \\ &\times \exp\left\{ t \left[ \exp\left( \frac{it}{\sqrt{v}} \right) - 1 \right] \right\} \left[ \int_0^v \exp\left\{ \left[ 1 - \exp\left( \frac{it}{\sqrt{v}} \right) \right] u \right\} B_{\exp(it/\sqrt{v})}(u) du \right] dt. \end{aligned}$$

Further, we obtain

$$\begin{aligned} (18) \quad &|\sqrt{v} P_k(v) - \sqrt{v} A_k(v)| \\ &\leq \frac{1}{2\pi} \int_{-\pi\sqrt{v}}^{\pi\sqrt{v}} \left| \exp\left\{ \frac{it}{\sqrt{v}} \right\} - 1 \right| \left| \int_0^v \exp\left\{ \left[ \exp\left( \frac{it}{\sqrt{v}} \right) - 1 \right] (v-u) \right\} \varrho(u) du \right| dt \\ &\leq \frac{1}{\pi\sqrt{v}} \int_0^v \varrho(u) \left[ \int_0^{\pi\sqrt{v}} t \exp\left\{ -2(v-u) \sin^2 \frac{t}{2\sqrt{v}} \right\} dt \right] du \\ &= \frac{1}{\pi\sqrt{v}} \int_0^v \varrho(u) \left[ \int_0^{\pi\sqrt{v}\varepsilon} t \exp\left\{ -2(v-u) \sin^2 \frac{t}{2\sqrt{v}} \right\} dt \right] du + \\ &\quad + \frac{1}{\pi\sqrt{v}} \int_0^v \varrho(u) \left[ \int_{\pi\sqrt{v}\varepsilon}^{\pi\sqrt{v}} t \exp\left\{ -2(v-u) \sin^2 \frac{t}{2\sqrt{v}} \right\} dt \right] du, \quad 0 < \varepsilon < 1. \end{aligned}$$

We now estimate the first integral in (18). Since  $0 < t/2\sqrt{v} < \pi\varepsilon/2$ , choosing  $\varepsilon > 0$  sufficiently small we can find  $\alpha > 0$  such that  $\sin(t/2\sqrt{v}) > \alpha t/2\sqrt{v}$ . Therefore, we obtain

$$\begin{aligned} &\frac{1}{\pi\sqrt{v}} \int_0^v \varrho(u) \left[ \int_0^{\pi\sqrt{v}\varepsilon} t \exp\left\{ -2(v-u) \sin^2 \frac{t}{2\sqrt{v}} \right\} dt \right] du \\ &\leq \frac{1}{\pi\sqrt{v}} \int_0^v \varrho(u) \left[ \int_0^{\pi\sqrt{v}\varepsilon} t \exp\left\{ -\alpha \frac{t^2}{2v} (v-u) \right\} dt \right] du \\ &= \frac{1}{\alpha\pi\sqrt{v}} \int_0^v \frac{v}{v-u} \varrho(u) \left( 1 - \exp\left\{ -\frac{\alpha\pi^2\varepsilon^2(v-u)}{2} \right\} \right) du \\ &= \frac{1}{\alpha\pi\sqrt{v}} \int_0^v \varrho(u) \left( 1 - \exp\left\{ -\frac{\alpha\pi^2\varepsilon^2(v-u)}{2} \right\} \right) du + \end{aligned}$$



$$\begin{aligned}
& + \frac{1}{\alpha\pi\sqrt{v}} \int_0^v \frac{u}{v-u} \varrho(u) \left( 1 - \exp \left\{ -\frac{\alpha\pi^2 \varepsilon^2 (v-u)}{2} \right\} \right) du \\
& \leq \frac{c_1}{\sqrt{v}} \int_0^v \varrho(u) du + \frac{c_2}{\sqrt{v}} \int_0^v u \varrho(u) du, \quad 0 < c_1, c_2 < \infty.
\end{aligned}$$

Hence the first integral in (18) tends to zero as  $v \rightarrow \infty$ .

We now estimate the second integral in (18). Clearly, we have

$$\exp \left\{ -2 \sin^2 \frac{t}{2\sqrt{v}} \right\} < e^{-c}, \quad \pi\sqrt{v}\varepsilon < t < \pi\sqrt{v}, \quad 0 < \varepsilon < 1, \quad 0 < c < \infty.$$

Therefore, we obtain

$$\begin{aligned}
& \frac{1}{\pi\sqrt{v}} \int_0^v \varrho(u) \left[ \int_{\pi\sqrt{v}\varepsilon}^{\pi\sqrt{v}} t \exp \left\{ -2(v-u) \sin^2 \frac{t}{2\sqrt{v}} \right\} dt \right] du \\
& \leq \frac{1}{\pi\sqrt{v}} \int_0^v \varrho(u) \left[ \int_{\pi\sqrt{v}\varepsilon}^{\pi\sqrt{v}} t e^{-c(v-u)} dt \right] du = \frac{c_3 v}{\pi\sqrt{v}} \int_0^v e^{-c(v-u)} \varrho(u) du \\
& = c_4 \sqrt{v} e^{-cv} \int_0^{v\gamma} e^{cu} \varrho(u) du + c_4 \sqrt{v} e^{-cv} \int_{v\gamma}^v e^{cu} \varrho(u) du, \\
& \qquad \qquad \qquad 0 < c_3, c_4 < \infty, \quad 0 < \gamma < 1.
\end{aligned}$$

We complete the proof of the theorem observing that

$$\begin{aligned}
\sqrt{v} e^{-cv} \int_0^{v\gamma} e^{cu} \varrho(u) du & \leq \sqrt{v} e^{-cv} e^{cv\gamma} \int_0^{v\gamma} \varrho(u) du \\
& = v e^{-cv(1-\gamma)} \frac{1}{\sqrt{v}} \int_0^{v\gamma} \varrho(u) du \rightarrow 0 \quad \text{as } v \rightarrow \infty
\end{aligned}$$

and

$$\begin{aligned}
\sqrt{v} e^{-cv} \int_{v\gamma}^v e^{cu} \varrho(u) du & \leq \sqrt{v} \int_{v\gamma}^v \varrho(u) du \leq \sqrt{v} \int_{v\gamma}^v \frac{u}{v\gamma} \varrho(u) du \\
& \leq \frac{1}{\gamma\sqrt{v}} \int_0^v u \varrho(u) du \rightarrow 0 \quad \text{as } v \rightarrow \infty.
\end{aligned}$$

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